# Instrumental Variables<sup>1</sup> STA431 Winter/Spring 2013

<sup>&</sup>lt;sup>1</sup>See last slide for copyright information.







## Seeking identifiability

We have seen that in simple regression, parameters of a model with measurement error are not identifiable.

$$Y_i = \alpha_1 + \beta_1 X_i + \epsilon_i$$
  
$$W_i = \nu + X_i + e_i,$$

- For example, X might be income and Y might be credit card debt.
- Include another response variable  $Y_2$ , like value of automobile.

### Include a second response variable

- Response variable of primary interest is now called  $Y_{i,1}$
- The second response variable  $Y_{i,2}$  is called an *instrumental* variable.
- It's just a tool.

$$W_i = \nu + X_i + e_i$$
  

$$Y_{i,1} = \alpha_1 + \beta_1 X_i + \epsilon_{i,1}$$
  

$$Y_{i,2} = \alpha_2 + \beta_2 X_i + \epsilon_{i,2}$$

where  $X_i$ ,  $e_i$ ,  $\epsilon_{i,1}$  and  $\epsilon_{i,2}$  are all independent,  $Var(X_i) = \phi$ ,  $Var(e_i) = \omega$ ,  $Var(\epsilon_{i,1}) = \psi_1$ ,  $Var(\epsilon_{i,2}) = \psi_2$ ,  $E(X_i) = \mu_x$  and the expected values of all error terms are zero. The regression coefficients  $\alpha_j$  and  $\beta_j$  are unknown constants.

## Are the parameters identifiable?

$$\begin{array}{rcl} W_{i} & = & \nu + X_{i} + e_{i} \\ Y_{i,1} & = & \alpha_{1} + \beta_{1}X_{i} + \epsilon_{i,1} \\ Y_{i,2} & = & \alpha_{2} + \beta_{2}X_{i} + \epsilon_{i,2} \end{array}$$

- Assume everything is normal:  $\mathbf{D}_i \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ .
- $\boldsymbol{\theta} = (\nu, \alpha_1, \alpha_2, \beta_1, \beta_2, \mu_x, \phi, \omega, \psi_1, \psi_2)$ : Ten parameters.
- $\boldsymbol{\mu}$  is  $3 \times 1$ .
- $\Sigma$  is  $3 \times 3$ .
- Nine moment structure equations in ten unknowns.

•

### Look at the covariance structure equations We are pessimistic about the expected values

$$W_i = \nu + X_i + e_i$$
  

$$Y_{i,1} = \alpha_1 + \beta_1 X_i + \epsilon_{i,1}$$
  

$$Y_{i,2} = \alpha_2 + \beta_2 X_i + \epsilon_{i,2}$$

$$\boldsymbol{\Sigma} = V \begin{pmatrix} W_i \\ Y_{i,1} \\ Y_{i,2} \end{pmatrix} = \begin{bmatrix} \phi + \omega & \beta_1 \phi & \beta_2 \phi \\ & \beta_1^2 \phi + \psi_1 & \beta_1 \beta_2 \phi \\ & & & \beta_2^2 \phi + \psi_2 \end{bmatrix}$$

### Six equations in six unknowns A unique solution is possible but not guaranteed

$$\begin{bmatrix} \sigma_{11} & \sigma_{12} & \sigma_{13} \\ \sigma_{22} & \sigma_{23} \\ \sigma_{33} \end{bmatrix} = \begin{bmatrix} \phi + \omega & \beta_1 \phi & \beta_2 \phi \\ & \beta_1^2 \phi + \psi_1 & \beta_1 \beta_2 \phi \\ & & & \beta_2^2 \phi + \psi_2 \end{bmatrix}$$

Identifiability depends on where you are in the parameter space. Consider

- $\beta_1 = 0$  and  $\beta_2 = 0$
- $\beta_1 = 0$  and  $\beta_2 \neq 0$
- $\beta_1 \neq 0$  and  $\beta_2 = 0$
- $\beta_1 \neq 0$  and  $\beta_2 \neq 0$

The parameter  $\beta_1$  is identifiable if  $\beta_2 \neq 0$ :  $\beta_1 = \frac{\sigma_{23}}{\sigma_{13}}$ .

## Suppose both $\beta_1 \neq 0$ and $\beta_2 \neq 0$

$$\begin{bmatrix} \sigma_{11} & \sigma_{12} & \sigma_{13} \\ \sigma_{22} & \sigma_{23} \\ \sigma_{33} \end{bmatrix} = \begin{bmatrix} \phi + \omega & \beta_1 \phi & \beta_2 \phi \\ & \beta_1^2 \phi + \psi_1 & \beta_1 \beta_2 \phi \\ & & \beta_2^2 \phi + \psi_2 \end{bmatrix}$$

$$\beta_1 = \frac{\sigma_{23}}{\sigma_{13}}$$
$$\beta_2 = \frac{\sigma_{23}}{\sigma_{12}}$$
$$\phi = \frac{\sigma_{12}\sigma_{13}}{\sigma_{23}}$$

Solve for  $\omega$ ,  $\psi_1$  and  $\psi_2$  by subtraction. Can write

$$\begin{split} & \omega = \sigma_{11} - \phi \\ & \psi_1 = \sigma_{22} - \beta_1^2 \phi \\ & \psi_2 = \sigma_{33} - \beta_2^2 \phi \end{split}$$

Without substituting for parameter that have already been identified. Don't need to give complete explicit solution. This shows it can be done.

### What about the expected values?

$$\begin{array}{rcl} W_{i} & = & \nu + X_{i} + e_{i} \\ Y_{i,1} & = & \alpha_{1} + \beta_{1} X_{i} + \epsilon_{i,1} \\ Y_{i,2} & = & \alpha_{2} + \beta_{2} X_{i} + \epsilon_{i,2} \end{array}$$

$$\mu_1 = \nu + \mu_x$$
  

$$\mu_2 = \alpha_1 + \beta_1 \mu_x$$
  

$$\mu_3 = \alpha_2 + \beta_2 \mu_x$$

- Three equations in four unknowns, even assuming  $\beta_1$  and  $\beta_2$  known.
- Re-parameterize.

### Re-parameterize

$$\mu_1 = \nu + \mu_x$$
  

$$\mu_2 = \alpha_1 + \beta_1 \mu_x$$
  

$$\mu_3 = \alpha_2 + \beta_2 \mu_x$$

- Absorb  $\nu, \mu_x, \alpha_1, \alpha_2$  into  $\mu$ .
- Parameter was  $\boldsymbol{\theta} = (\nu, \mu_x, \alpha_1, \alpha_2, \beta_1, \beta_2, \phi, \omega, \psi_1, \psi_2)$
- Now it's  $\boldsymbol{\theta} = (\mu_1, \mu_2, \mu_3, \beta_1, \beta_2, \phi, \omega, \psi_1, \psi_2)$
- Dimension of the parameter space is now one less.
- We haven't lost much.

One Explanatory Variable

Multiple Explanatory Variables

# We haven't lost much especially because the model was already re-parameterized

Model is

$$\begin{split} W_i &= \nu + X_i + e_i \\ Y_{i,1} &= \alpha_1 + \beta_1 X_i + \epsilon_{i,1} \\ Y_{i,2} &= \alpha_2 + \beta_2 X_i + \epsilon_{i,2} \end{split}$$

But of course there is measurement error in  $Y_1$  and  $Y_2$ . Recall

$$Y = \alpha + \beta X + \epsilon$$
  

$$V = \nu_0 + Y + e$$
  

$$= \nu_0 + (\alpha + \beta X + \epsilon) + e$$
  

$$= (\nu_0 + \alpha) + \beta X + (\epsilon + e)$$
  

$$= \alpha' + \beta X + \epsilon'$$

### Summary

- Adding the instrumental variable didn't help identify the expected values and intercepts. That's hopeless.
- But we did identify  $\beta_1$ , which is the most interesting parameter.
- Re-parameterizing, absorbed the intercepts and expected values into  $\mu$ .
- Where  $\beta_1$  and  $\beta_2$  are both non-zero, the entire parameter vector is identifiable.
- For maximum likelihood estimation, it helps to have the *entire* parameter vector identifiable at the true parameter value.
- This is definitely a success.

### Testing $H_0: \beta_1 = 0$ The most interesting null hypothesis

- The parameter  $\beta_1$  is identifiable, so a valid test is possible.
- But the whole parameter *vector* is not identifiable when  $\beta_1 = 0$ .
- Technical conditions of the likelihood ratio test are not satisfied.
- It becomes quite "interesting."
- Likelihood ratio statistic actually has 2 df even though  $H_0$  appears to impose only one restriction on the parameter.
- Too interesting.

•

## It's better with two (or more) instrumental variables.

$$\begin{split} W_i &= \nu + X_i + e_i \\ Y_{i,1} &= \alpha_1 + \beta_1 X_i + \epsilon_{i,1} \\ Y_{i,2} &= \alpha_2 + \beta_2 X_i + \epsilon_{i,2} \\ Y_{i,3} &= \alpha_3 + \beta_3 X_i + \epsilon_{i,3}, \end{split}$$

## With two instrumental variables

- Again, identification of the expected values and intercepts is out of the question.
- So we re-parameterize,
- Absorbing the expected values and intercepts into µ = E(D<sub>i</sub>)
- And look at the covariance structure equations.

### Covariance structure equations

- Ten equations in eight unknowns.
- Unique solution possible but not guaranteed.
- Primary interest is still in  $\beta_1$ .
- Assume  $\beta_2 \neq 0$  and  $\beta_3 \neq 0$ , meaning only that  $Y_2$  and  $Y_3$  are well chosen.

One Explanatory Variable

Multiple Explanatory Variables

Solve for  $\phi$ Assuming  $\beta_2 \neq 0$  and  $\beta_3 \neq 0$ 

$$\begin{pmatrix} \sigma_{11} & \sigma_{12} & \sigma_{13} & \sigma_{14} \\ \sigma_{22} & \sigma_{23} & \sigma_{24} \\ \sigma_{33} & \sigma_{3,4} \\ & & & \sigma_{4,4} \end{pmatrix} = \begin{pmatrix} \phi + \omega & \beta_1 \phi & \beta_2 \phi & \beta_3 \phi \\ & \beta_1^2 \phi + \psi_1 & \beta_1 \beta_2 \phi & \beta_1 \beta_3 \phi \\ & & & \beta_2^2 \phi + \psi_2 & \beta_2 \beta_3 \phi \\ & & & & & \beta_3^2 \phi + \psi_3 \end{pmatrix}$$

$$\frac{\sigma_{13}\sigma_{14}}{\sigma_{34}} = \frac{\beta_2\beta_3\phi^2}{\beta_2\beta_3\phi} = \phi$$

### Then all you have to write is

$$\begin{split} \omega &= \sigma_{11} - \phi \\ \beta_1 &= \frac{\sigma_{12}}{\phi} \\ \beta_2 &= \frac{\sigma_{13}}{\phi} \\ \beta_3 &= \frac{\sigma_{14}}{\phi} \\ \psi_1 &= \sigma_{22} - \beta_1^2 \phi \\ \psi_2 &= \sigma_{33} - \beta_2^2 \phi \\ \psi_3 &= \sigma_{44} - \beta_3^2 \phi \end{split}$$

Notice again how once we have solved for a model parameter, we may use it to solve for other parameters without explicitly substituting in terms of  $\sigma_{ij}$ .

### Parameters can be recovered from the covariance matrix

$$\phi = \frac{\sigma_{13}\sigma_{14}}{\sigma_{34}} \qquad \beta_3 = \frac{\sigma_{14}}{\phi}$$

$$\omega = \sigma_{11} - \phi \qquad \psi_1 = \sigma_{22} - \beta_1^2 \phi$$

$$\beta_1 = \frac{\sigma_{12}}{\phi} \qquad \psi_2 = \sigma_{33} - \beta_2^2 \phi$$

$$\beta_2 = \frac{\sigma_{13}}{\phi} \qquad \psi_3 = \sigma_{44} - \beta_3^2 \phi$$

- Parameter vector is identifiable almost everywhere in the parameter space.
- Everywhere  $\beta_2$  and  $\beta_3$  are both non-zero
- $\beta_1 = 0 \Leftrightarrow \sigma_{12} = 0$  presents no problem.

# But there is more than one way to recover the parameter values from $\Sigma$

$$\begin{pmatrix} \sigma_{11} & \sigma_{12} & \sigma_{13} & \sigma_{14} \\ \sigma_{22} & \sigma_{23} & \sigma_{24} \\ \sigma_{33} & \sigma_{3,4} \\ & & & & & & & \\ \end{pmatrix} = \begin{pmatrix} \phi + \omega & \beta_1 \phi & \beta_2 \phi & \beta_3 \phi \\ & \beta_1^2 \phi + \psi_1 & \beta_1 \beta_2 \phi & \beta_1 \beta_3 \phi \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & &$$

$$\beta_1 = \frac{\sigma_{12}\sigma_{34}}{\sigma_{13}\sigma_{14}}$$
$$\beta_1 = \frac{\sigma_{23}}{\sigma_{13}}$$
$$\beta_1 = \frac{\sigma_{24}}{\sigma_{14}}$$

## Is there a problem?

$$\beta_1 = \frac{\sigma_{12}\sigma_{34}}{\sigma_{13}\sigma_{14}} = \frac{\sigma_{23}}{\sigma_{13}} = \frac{\sigma_{24}}{\sigma_{14}}$$

Does this mean the solution for  $\beta_1$  is not "unique?"

- No everything is okay.
- If the parameters can be recovered from the covariances in any way at all, they are identifiable.
- If the model is correct, all the seemingly different ways must be the same.
- That is,

$$\frac{\sigma_{12}\sigma_{34}}{\sigma_{13}\sigma_{14}} = \frac{\sigma_{23}}{\sigma_{13}} \text{ and } \frac{\sigma_{12}\sigma_{34}}{\sigma_{13}\sigma_{14}} = \frac{\sigma_{24}}{\sigma_{14}}$$

• Simplifying a bit,

$$\sigma_{12}\sigma_{34} = \sigma_{14}\sigma_{23} = \sigma_{13}\sigma_{24}$$

### Model implies two constraints on the covariance matrix

## $\sigma_{12}\sigma_{34} = \sigma_{14}\sigma_{23} = \sigma_{13}\sigma_{24}$

- All three products equal  $\beta_1\beta_2\beta_3\phi^2$
- True even when some  $\beta_j = 0$

$$\begin{pmatrix} \sigma_{11} & \sigma_{12} & \sigma_{13} & \sigma_{14} \\ \sigma_{22} & \sigma_{23} & \sigma_{24} \\ \sigma_{33} & \sigma_{3,4} \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & & \\ & & & & & & & & & & \\ & & & & & & & & & & & \\ & & & & & & & & & & & \\ & & & & & & & & & & & \\ & & & & & & & & & & & \\ & & & & & & & & & & & \\ & & & & & & & & & & & \\ & & & & & & & & & & & \\ & & & & & & & & & & & \\ & & & & & & & & & & \\ & & & & & & & & & & \\ & & & & & & & & & & \\ & & & & & & & & & & \\ & & & & & & & & & & \\ & & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & &$$

## Model implies $\sigma_{12}\sigma_{34} = \sigma_{14}\sigma_{23} = \sigma_{13}\sigma_{24}$

- Parameter is identifiable.
- Ten equations in eight unknowns.
- Call the parameter *over-identifiable*.
- If there were the same number of equations as unknowns, it would be *just identifiable*.
- Model imposes two equality constraints (restrictions) on the covariance matrix: 10 8 = 2
- Sometimes called *over-identifying restrictions*.
- These are the constraints that are tested in the likelihood ratio test for goodness of fit.
- More instrumental variables can't hurt.

### Multiple Explanatory Variables An example with just two explanatory variables (and two instrumental variables)

Independently for  $i = 1, \ldots, n$ ,

$$W_{i,1} = \nu_1 + X_{i,1} + e_{i,1}$$

$$Y_{i,1} = \alpha_1 + \beta_1 X_{i,1} + \epsilon_{i,1}$$

$$Y_{i,2} = \alpha_2 + \beta_2 X_{i,1} + \epsilon_{i,2}$$

$$W_{i,2} = \nu_2 + X_{i,2} + e_{i,2}$$

$$Y_{i,3} = \alpha_3 + \beta_3 X_{i,2} + \epsilon_{i,3}$$

$$Y_{i,4} = \alpha_4 + \beta_4 X_{i,2} + \epsilon_{i,4}$$

where  $E(X_{i,j}) = \mu_j$ ,  $e_{i,j}$  and  $\epsilon_{i,j}$  are independent of one another and of  $X_{i,j}$ ,  $Var(e_{i,j}) = \omega_j$ ,  $Var(\epsilon_{i,j}) = \psi_j$ , and

$$V\left(\begin{array}{c}X_{i,1}\\X_{i,1}\end{array}\right) = \left(\begin{array}{c}\phi_{11} & \phi_{12}\\\phi_{12} & \phi_{22}\end{array}\right)$$

# As usual, intercepts and expected values can't be recovered individually

- Eight intercepts and expected values of latent variables.
- Six expected values of observable variables.
- Re-parameterize, absorbing them into  $\mu_1, \ldots, \mu_6$ .
- Estimate with the vector of 6 sample means and set them aside, forever.

# Covariance matrix of $(W_{i,1}, Y_{i,1}, Y_{i,2}, W_{i,2}, Y_{i,3}, Y_{i,4})'$

$$[\sigma_{ij}] =$$



$$\boldsymbol{\theta} = (\beta_1, \beta_2, \beta_3, \beta_4, \phi_{11}, \phi_{12}, \phi_{22}, \omega_1, \omega_2, \psi_1, \psi_2, \psi_3, \psi_4)$$

- Does this model pass the test of the parameter count rule?
- Where the parameter vector is identifiable, how many over-identifying restrictions are there?
- How many degrees of freedom in the likelihood ratio test for model fit?

### Where is the entire parmeter vector identifiable?



- What happens if  $\beta_1 = \beta_2 = 0$ ?
- Why is it reasonable to assume  $\beta_2 \neq 0$  and  $\beta_4 \neq 0$ ?
- In that case, what else do you need?
- Would any other condition identify the whole parameter vector?

### My answer



#### EITHER

- One of  $\beta_1$  and  $\beta_2$  non-zero, and
- One of  $\beta_3$  and  $\beta_4$  non-zero, and
- $\phi_{12} \neq 0$

OR, all of  $\beta_1, \ldots, \beta_4$  non-zero.

### Could we get by with less information? If we wanted to identify just some interesting parameters?



- Usual rule in Econometrics is at least one instrumental variable for each explanatory variable.
- What if no instrumental variable for  $X_2$ ?
- What if no response variables at all for  $X_2$ ?

## Observations

- Instrumental variables can solve some of the terrible problems with measurement error in regression.
- General rules like "At least one instrumental variable for each explanatory variable" are useful even if they are over-simplifications.
- Awareness of parameter identifiability is vital in the *planning* of data collection.
- Most observational data sets are collected without the right kind of planning.

# Copyright Information

This slide show was prepared by Jerry Brunner, Department of Statistical Sciences, University of Toronto. It is licensed under a Creative Commons Attribution - ShareAlike 3.0 Unported License. Use any part of it as you like and share the result freely. The LATEX source code is available from the course website:

http://www.utstat.toronto.edu/~brunner/oldclass/431s31