

# Assignment 6

[1]

$$\begin{aligned} \textcircled{1} \quad P(T > t_j | T > t_{j-1}) &= \frac{P(T > t_j \cap T > t_{j-1})}{P(T > t_{j-1})} \\ &= \frac{P(T > t_j)}{P(T > t_{j-1})} = \frac{S(t_j)}{S(t_{j-1})} \end{aligned}$$

$$\begin{aligned} \textcircled{2} \quad \prod_{j=1}^k P_j &= \prod_{j=1}^k \frac{S(t_j)}{S(t_{j-1})} \\ &= \frac{\cancel{S(t_1)}}{\cancel{S(t_0)}} \cdot \frac{\cancel{S(t_2)}}{\cancel{S(t_1)}} \cdot \frac{\cancel{S(t_3)}}{\cancel{S(t_2)}} \cdots \frac{\cancel{S(t_{k-1})}}{\cancel{S(t_{k-2})}} \cdot \frac{S(t_k)}{\cancel{S(t_{k-1})}} \\ &= S(t_k) \end{aligned}$$

$$\textcircled{3} \quad (\text{a}) \quad E(\hat{p}) = p, \quad \text{Var}(\hat{p}) = \frac{p(1-p)}{n} \quad \left( = \frac{\sigma^2}{n} \right)$$

$$(\text{b}) \quad \hat{p} \sim N(p, \frac{p(1-p)}{n})$$

$$\begin{aligned} \textcircled{4} \quad \hat{p}_j &= \cancel{d_j} \mid - \frac{d_j}{n_j} \quad \left( \begin{array}{l} \text{one minus conditional} \\ \text{probability of death} \end{array} \right) \\ &= \frac{n_j - d_j}{n_j} \quad \text{The proportion who lived} \end{aligned}$$

$$\textcircled{5} \quad (\text{a}) \quad \hat{p}_j \sim N(p_j, \frac{p_j(1-p_j)}{n_j})$$

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(5b) Using the one-variable delta method,  $\hat{P}_j$  should be asymptotically normal with expected value  $P_j$ . To get the variance, with  $g(x) = \log x$

$$g'(x) = \frac{d}{dx} \log x = \frac{1}{x}, \text{ so asymptotic variance is}$$

$$\frac{1}{P_j^2} \frac{P_j(1-P_j)}{n_j} = \frac{1-P_j}{n_j P_j}, \text{ and (loosely)}$$

$$\log \hat{P}_j \sim N\left(\log P_j, \frac{1-P_j}{n_j P_j}\right)$$

⑥ (a)  $\log \hat{S}(t) = \sum_{t_j \leq t} \log \hat{P}_j$

(b)  $E(\log \hat{S}(t)) \approx \sum_{t_j \leq t} E(\log \hat{P}_j)$   
 $= \sum_{t_j \leq t} \log P_j$

(c)  $V_m(\log \hat{S}(t)) \approx \sum_{t_j \leq t} V_m(\log \hat{P}_j)$   
 $= \sum_{t_j \leq t} \frac{1-P_j}{n_j P_j}$  so we guess

or  $\log(S(A))$

(d)  $\log \hat{S}(t) \sim N\left(\underbrace{\sum_{t_j \leq t} \log P_j}_{\text{or } \log(S(A))}, \sum_{t_j \leq t} \frac{1-P_j}{n_j P_j}\right)$

3

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Have asymptotic distribution of  $\log \hat{S}(t)$ . Want distribution of  $g(\log \hat{S}(t)) = \exp \{\log \hat{S}(t)\}$ . Using one-variable delta method  $g'(x) = g(x)$ , so

$$\hat{S}(t) \sim N(S(t), S(t)^2 \sum_{t_j \leq t} \frac{1 - \hat{p}_j}{n_j \hat{p}_j})$$

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Standard error is the square root of the estimated variance. Estimated variance  $\rightarrow$

$$\begin{aligned} & \hat{S}(t)^2 \sum_{t_j \leq t} \frac{1 - \hat{p}_j}{n_j \hat{p}_j} = \hat{S}(t)^2 \sum_{t_j \leq t} \frac{d_j / n_j}{n_j / n_j - d_j / n_j} \\ & = \hat{S}^2(t) \sum_{t_j \leq t} \frac{d_j}{n_j(n_j - d_j)} \quad , \text{ and standard error is} \end{aligned}$$

$$\hat{S}(t) \sqrt{\sum_{t_j \leq t} \frac{d_j}{n_j(n_j - d_j)}}, \text{ where}$$

$$\hat{S}(t) = \prod_{t_j \leq t} \hat{p}_j = \prod_{t_j \leq t} \left(1 - \frac{d_j}{n_j}\right)$$

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$t_j$	$n_j$	$d_j$	$\hat{p}_j$	$\hat{S}(t_j)$
0	100	0	1	1
2	100	15	$85/100$	0.85
4	83	5	$78/83$	0.7988
5	73	10	$63/73$	0.6894

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5 observations were censored  
 between  $t=4$  &  $t=5$

5

- 12 (a) For an exponential distribution get median

by setting  $S(\lambda) = \frac{1}{2}$ , so  $e^{-\lambda t} = \frac{1}{2} \Rightarrow -\lambda t = \log \frac{1}{2}$   
 $-\lambda t = \log(2^{-1}) = -\log 2 \Rightarrow \lambda t = \log 2 \Rightarrow t = \frac{\log 2}{\lambda}$

And MLE of median is  $\frac{\log 2}{\lambda}$

To get a confidence interval for the median, either use the delta method or transform the confidence interval for  $\lambda$ .

- Delta method. From Question 1b of assignment 5,

$$\text{Var}(\hat{\lambda}) \approx \frac{\lambda^2}{\sum s_i} = N_n \cdot g(\lambda) = (\log 2) \lambda^{-1}$$

$$g'(\lambda) = -(\log 2) \lambda^{-2} = -\frac{108 2}{\lambda^2}, \text{ so asymptotic}$$

$$\text{variance of median is } g'(\lambda)^2 N_n = \frac{(\log 2)^2}{\lambda^4} \frac{\lambda^2}{\sum s_i}$$

$$= \frac{(\log 2)^2}{\lambda^2 \sum s_i}, \text{ and standard error is}$$

$$S_{\text{med}} = \frac{\log 2}{\sqrt{\lambda \sum s_i}}$$

$$S_{\text{med}} = \frac{\log 2}{\sqrt{\lambda \sum s_i}}$$

- Transform confidence interval.

$$0.95 \approx P(A < \lambda < B) = P\left(\frac{1}{A} > \frac{1}{\lambda} > \frac{1}{B}\right)$$

$$= P\left(\frac{\log 2}{B} < \frac{\log 2}{\lambda} = \text{med} < \frac{\log 2}{A}\right)$$

```
R version 4.2.2 (2022-10-31) -- "Innocent and Trusting"
Copyright (C) 2022 The R Foundation for Statistical Computing
Platform: x86_64-apple-darwin17.0 (64-bit)
```

```
R is free software and comes with ABSOLUTELY NO WARRANTY.
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Type 'license()' or 'licence()' for distribution details.
```

```
Natural language support but running in an English locale
```

```
R is a collaborative project with many contributors.
Type 'contributors()' for more information and
'citation()' on how to cite R or R packages in publications.
```

```
Type 'demo()' for some demos, 'help()' for on-line help, or
'help.start()' for an HTML browser interface to help.
Type 'q()' to quit R.
```

```
[R.app GUI 1.79 (8160) x86_64-apple-darwin17.0]
```

```
[Workspace restored from /Users/brunner/.RData]
[History restored from /Users/brunner/.Rapp.history]
```

```
> # A6
>
> rm(list=ls()); options(scipen=999)
> exdata = read.table("http://www.utstat.utoronto.ca/brunner/data/legal/expo.data2.txt")
> head(exdata); Time = exdata$Time; Uncensored = exdata$Uncensored
  Time Uncensored
1 0.179      0
2 1.024      1
3 0.189      1
4 0.345      1
5 0.977      1
6 0.241      1
>
> # 11(a)
> # install.packages("survival",dependencies=TRUE) # Only need to do this once
> library(survival)
> y = Surv(Time,Uncensored)
> km = survfit(y ~ 1); km
Call: survfit(formula = y ~ 1)

      n events median 0.95LCL 0.95UCL
[1,] 50      40  0.351   0.284   0.758
> sumkm = summary(km); sumkm
Call: survfit(formula = y ~ 1)

    time n.risk n.event survival std.err lower 95% CI upper 95% CI
0.026    50      2  0.9600  0.0277   0.90719  1.000
0.032    47      1  0.9396  0.0338   0.87557  1.000
0.058    44      1  0.9182  0.0392   0.84448  0.998
0.062    43      1  0.8969  0.0437   0.81511  0.987
0.100    41      1  0.8750  0.0478   0.78610  0.974
0.101    40      1  0.8531  0.0514   0.75811  0.960
0.109    39      1  0.8312  0.0545   0.73095  0.945
0.117    38      1  0.8094  0.0573   0.70448  0.930
0.118    37      1  0.7875  0.0598   0.67860  0.914
0.165    36      1  0.7656  0.0620   0.65324  0.897
0.173    35      1  0.7437  0.0640   0.62835  0.880
0.179    34      1  0.7219  0.0657   0.60388  0.863
0.189    32      1  0.6993  0.0674   0.57888  0.845
0.239    31      1  0.6768  0.0689   0.55428  0.826
0.241    30      1  0.6542  0.0702   0.53007  0.807
```

0.265	29	1	0.6316	0.0713	0.50621	0.788
0.284	28	1	0.6091	0.0723	0.48270	0.769
0.318	27	1	0.5865	0.0730	0.45951	0.749
0.338	26	1	0.5640	0.0736	0.43665	0.728
0.345	25	1	0.5414	0.0741	0.41409	0.708
0.350	24	1	0.5188	0.0743	0.39184	0.687
0.351	23	1	0.4963	0.0744	0.36988	0.666
0.450	21	1	0.4727	0.0745	0.34697	0.644
0.466	20	1	0.4490	0.0745	0.32441	0.622
0.478	19	1	0.4254	0.0742	0.30220	0.599
0.499	18	1	0.4018	0.0738	0.28035	0.576
0.514	17	1	0.3781	0.0731	0.25886	0.552
0.515	16	1	0.3545	0.0723	0.23774	0.529
0.634	15	1	0.3309	0.0712	0.21701	0.504
0.758	13	1	0.3054	0.0701	0.19473	0.479
0.864	10	1	0.2749	0.0694	0.16752	0.451
0.977	8	1	0.2405	0.0687	0.13736	0.421
1.024	7	1	0.2061	0.0670	0.10907	0.390
1.027	6	1	0.1718	0.0640	0.08277	0.357
1.068	5	1	0.1374	0.0597	0.05864	0.322
1.172	4	1	0.1031	0.0538	0.03708	0.287
1.188	3	1	0.0687	0.0455	0.01876	0.252
1.601	2	1	0.0344	0.0333	0.00514	0.230
1.836	1	1	0.0000	NaN	NA	NA

```

>
> # (b) S-hat(0:062) = 0.8969
>
> # (c) and (d)
> # p-hat      1      2      3      4
> #           48/50 * 46/47 * 43/44 * 42/43
[1] 0.8968665
>
>
> # (e) This way of doing it requires you to realize sumkm is a list.
> #   The numbers could also be entered by hand.
> #   Try sumkm[1], sumkm[2] etc. to find out.
> n = sumkm$n.risk; d = sumkm$n.event; Shat = sumkm$surv
> Shat[4] # Another way to answer (b)
[1] 0.8968665
>
> se4 = Shat[4] * sqrt(sum(d[1:4]/(n[1:4]*(n[1:4]-d[1:4])))); se4 # 0.04373656
[1] 0.04373656
>
> # This should agree with the hand" calculation. Get 0.04373819
> 0.8969 * sqrt(sum(d[1:4]/(n[1:4]*(n[1:4]-d[1:4]))))
[1] 0.04373819
>
> # (f)
> plot(km)
>
>
> # 12
> # (a)
> # First, from last week (Assignment 5),
>
> # A5 Q2a) MLE
> lambdahat = sum(Uncensored)/sum(Time); lambdahat
[1] 1.717107
>
> # A5 Q2b Estimated asymptotic variance
> vhat = lambdahat^2 / sum(Uncensored); vhat # Estimated asymptotic variance
[1] 0.07371138
> se = sqrt(vhat); se
[1] 0.2714984
>
```

```

> # A5 Q2c) 95% CI for lambda: 0.95 =~ P(A < lambda < B)
> A = lambdahat - 1.96*se; B = lambdahat + 1.96*se
> c(A,B)
[1] 1.184970 2.249244
>
> # Now get to 12a from A6
> # Median of an exponential is log(2)/lambda
> medhat = log(2)/lambdahat; medhat
[1] 0.4036716
>
> # There are two ways to get a CI ...
>
> # Delta method
> semed = log(2)/(lambdahat*sqrt(sum(Uncensored))); semed
[1] 0.06382608
> lower95 = medhat - 1.96*semmed; upper95 = medhat + 1.96*semmed
> c(lower95,upper95)
[1] 0.2785725 0.5287707
>
> # Transform CI for lambda
> c(log(2)/B, log(2)/A)
[1] 0.3081690 0.5849492
>
>
>
> # (b) Add MLE of S(t) to plot
> t = seq(from=0,to=1.8,length=101)
> Shat = exp(-lambdahat*t)
> lines(t,Shat)
> title('Kaplan-Meier and MLE (MLE is smooth)')
>
>
```

## Kaplan-Meier and MLE (MLE is smooth)

