Moment-generating functions<sup>1</sup> STA 302: Fall 2014

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Let Y = g(X). There are two ways to get E(Y).

 $\bullet \quad \text{Derive the distribution of } Y \text{ and compute}$ 

$$E(Y) = \int_{-\infty}^{\infty} y \, f_Y(y) \, dy$$

**2** Use the distribution of X, and calculate

$$E(g(X)) = \int_{-\infty}^{\infty} g(x) \, f_X(x) \, dx$$

Big theorem: These two expressions are equal.

# The change of variables formula is very general Including but not limited to

$$E(g(X)) = \int_{-\infty}^{\infty} g(x) f_X(x) dx$$

$$E(g(\mathbf{X})) = \int_{-\infty}^{\infty} \cdots \int_{-\infty}^{\infty} g(x_1, \dots, x_p) f_{\mathbf{X}}(x_1, \dots, x_p) dx_1 \dots dx_p$$
$$E(g(X)) = \sum_x g(x) p_X(x)$$

## Moment-generating functions

$$M_{\scriptscriptstyle Y}(t) = E(e^{Yt}) = \left\{ \begin{array}{l} \int_{-\infty}^{\infty} e^{yt} \, f_{\scriptscriptstyle Y}(y) \, dy \\ \\ \sum_{y} e^{yt} p_{\scriptscriptstyle Y}(y) \end{array} \right.$$

## Properties of moment-generating functions

- Moment-generating functions can be used to generate moments. To get  $E(Y^k)$ , differentiate  $M_Y(t)$  with respect to t. Differentiate k times and set t = 0.
- Moment-generating functions correspond uniquely to probability distributions.

# The function M(t) is like a fingerprint of the probability distribution.

$$Y \sim N(\mu, \sigma^2)$$
 if and only if  $M_Y(t) = e^{\mu t + \frac{1}{2}\sigma^2 t^2}$ 

$$Y \sim \chi^2(\nu)$$
 if and only if  $M_{_Y}(t) = (1-2t)^{-\nu/2}$  for  $t < \frac{1}{2}$ 

# Normal: $M(t) = e^{\mu t + \frac{1}{2}\sigma^2 t^2}$

#### Fingerprints of the normal distribution



## Chi-squared: $M(t) = (1 - 2t)^{-\nu/2}$

#### Fingerprints of the chi-squared distribution



# Example: Using moment-generating functions to prove distribution facts

Let  $X \sim N(\mu, \sigma^2$ . Show  $Y = \frac{X - \mu}{\sigma} \sim N(0, 1)$ 

### Facts about moment-generating functions Use these to find distributions of *functions* of random variables

- $M_{aY}(t) = M_Y(at)$
- $M_{\scriptscriptstyle Y+a}(t) = e^{at} M_Y(t)$
- If  $Y_1, \ldots, Y_n$  are independent,  $M_{\sum_{i=1}^n Y_i}(t) = \prod_{i=1}^n M_{Y_i}(t)$

### If $W = W_1 + W_2$ with $W_1$ and $W_2$ independent, $W \sim \chi^2(\nu_1 + \nu_2)$ and $W_2 \sim \chi^2(\nu_2)$ then $W_1 \sim \chi^2(\nu_1)$ .

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http://www.utstat.toronto.edu/~brunner/oldclass/302f14