

# Probability and Stochastic Processes I I - Lecture 6e

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## VI.6 Continuous State Space Markov Chains

- here  $T = \mathbb{N}_0$  and  $\mathcal{S}$  with  $\sigma$ -algebra  $\mathcal{C}$  (could be  $\mathcal{S} = \mathbb{R}^k, \mathcal{C} = \mathcal{B}^k$ )

**Definition VI.11** A Markov kernel  $P : \mathcal{S} \times \mathcal{C} \rightarrow [0, 1]$  for a time homogeneous, Markov process with state space  $\mathcal{S}$  and initial probability measure  $\nu$  on  $(\mathcal{S}, \mathcal{C})$  satisfies (i)  $P(s, \cdot)$  is a probability measure on  $\mathcal{C}$  for every  $s \in \mathcal{S}$  and (ii)  $P(\cdot, C) : (\mathcal{S}, \mathcal{C}) \rightarrow (\mathbb{R}^1, \mathcal{B}^1)$  for every  $C \in \mathcal{C}$ .

- so  $P(\cdot, C)$  is a random variable and  $P(s, C)$  is interpreted as the conditional probability that the next state is in  $C$  given that the current state is  $s$

- note when  $\mathcal{S} \subset \mathbb{Z}$ , then  $P(\cdot, C)$  is given by  $P(i, \{j\}) = p_{ij}$

- so the time homogeneous Markov process with state space  $\mathcal{S}$  and Markov kernel  $P$  is given by  $\{X_n : n \in \mathbb{N}_0\}$  where  $X_0 \sim \nu$  and  $X_n | X_0, \dots, X_{n-1} \sim P(X_{n-1}, \cdot)$

- such Markov processes arise in many statistical problems where there is a (posterior) distribution on a continuous space we want to sample from but can only find a Markov chain Monte Carlo algorithm to do this approximately

### Example VI.11

-  $\{X_n : n \in \mathbb{N}_0\}$  is given by  $X_0 = 0$  ( $\nu$  is degenerate at 0) and  $P(x, \cdot) = N(x/2, 1)$  so  $X_1 \sim N(0, 1)$  and

$$\begin{aligned} P(X_2 \leq x_2) &= \int_{-\infty}^{\infty} P(X_2 \leq x_2 | X_1)(x_1) P_{X_1}(dx_1) \\ &= \int_{-\infty}^{\infty} \Phi\left(\frac{x_2 - x_1/2}{1}\right) \varphi(x_1) dx_1 \end{aligned}$$

but with  $Z \sim N(0, 1)$  independent of  $X_1$ , then  $X_2 = Z + X_1/2 \sim N(0, 1 + 1/4) = N(0, 5/4)$ ,  $X_3 = Z + X_2/2 \sim N(0, 1 + 5/16) = N(0, 21/16)$ , etc. ■

- note - as long as it is easy to generate from  $\nu$  and  $P(s, \cdot)$  for each  $s$ , then it is easy to simulate the process

- we need a generalization of the concept of irreducibility because in the continuous case  $P(s, \{x\}) = 0$  for all  $s, x$

- note

$$\begin{aligned}P^{(2)}(s, C) &= P(X_2 \in C \mid X_0 = s) \\&= \int_{\mathcal{S}} P(X_2 \in C \mid X_0 = s, X_1 = s') P(s, ds') \\&= \int_{\mathcal{S}} P(X_2 \in C \mid X_1 = s') P(s, ds') \text{ by MP} \\&= \int_{\mathcal{S}} P(s', C) P(s, ds') \text{ by TH}\end{aligned}$$

and  $P^{(n)}$  is obtained similarly

**Definition VI.12** A MC with state space  $\mathcal{S}$  is called  $\phi$ -irreducible, where  $\phi$  is a measure on  $\mathcal{C}$ , if for all  $s \in \mathcal{S}$ , whenever  $\phi(C) > 0$  then  $P^{(n)}(s, C) > 0$  for some  $n$ .

- typically  $\phi =$  volume measure and in the discrete case if  $\phi =$  counting measure this is the same definition in that case

**Proposition VI.35** If there is  $\delta > 0$  s.t.  $P(s, \cdot)$  has a positive density on  $[s - \delta, s + \delta]$  for every  $s \in \mathcal{S} = \mathbb{R}^1$ , then the MC with this kernel is  $\phi$ -irreducible where  $\phi$  is volume measure.

Proof: Suppose  $C \in \mathcal{B}^1$  and  $\phi(C) > 0$ . Then  $\exists m \in \mathbb{N}$  s.t.  $\phi(C \cap [-m, m]) > 0$ . Choose  $n$  s.t.  $s - n\delta < -m < m < s + n\delta$ . Then since  $P(s, \cdot)$  has positive density on  $[s - \delta, s + \delta]$  this implies  $P^{(n)}(s, \cdot)$  has positive density on  $[s - n\delta, s + n\delta] > 0$  (see \* below). But  $C \cap [-m, m] \subset [s - n\delta, s + n\delta]$  so

$$P^{(n)}(s, C) \geq P^{(n)}(s, C \cap [-m, m]) > 0.$$



\* if  $P(s, \cdot)$  has positive density on  $[s - \delta, s + \delta]$ , then so does  $P^{(2)}(s, \cdot)$  on  $[s - 2\delta, s + 2\delta]$  since

$$\begin{aligned} & P^{(2)}(s, [s - 2\delta, s + 2\delta]) \\ &= \int_{-\infty}^{\infty} P(x, [s - 2\delta, s + 2\delta]) P(s, dx) \\ &= \int_{-\infty}^{\infty} \left( \int_{s-2\delta}^{s+2\delta} f(z | x) dz \right) P(s, dx) \\ &\geq \int_{s-\delta}^{s+\delta} \left( \int_{s-2\delta}^{s+2\delta} f(z | x) dz \right) f(x | s) dx > 0 \end{aligned}$$

since  $f(z | x) > 0$  for all  $z \in [x - \delta, x + \delta]$  and this implies

$\int_{s-2\delta}^{s+2\delta} f(z | x) dz > 0$  whenever  $[x - \delta, x + \delta] \subset [s - 2\delta, s + 2\delta]$  and this holds whenever  $x \in [s - \delta, s + \delta]$  and note  $f(x | s) > 0$  there

**Definition VI.13** A MC with state space  $\mathcal{S}$  has *period*  $b$  if  $\mathcal{S}$  can be decomposed into mutually disjoint subsets  $\mathcal{S} = \mathcal{S}_0 \cup \dots \cup \mathcal{S}_{b-1}$  s.t.  $P(s, S_i) = 1$  whenever  $s \in S_{i-1}$  and  $i = [(i-1) + 1] \bmod b$ . If  $b = 1$ , then the MC is said to be *aperiodic*.

**Proposition VI.36** If there is  $\delta > 0$  s.t.  $P(s, \cdot)$  has a positive density on  $[s - \delta, s + \delta]$  for every  $s \in \mathcal{S} = \mathbb{R}^1$ , then a  $\phi$ -irreducible MC, where  $\phi$  is volume measure, is aperiodic.

Proof: Suppose  $b$  is the period and let  $s \in S_i$  where  $\text{vol}(S_i) > 0$ . Also, let  $m \in \mathbb{N}$  be s.t.  $\text{vol}(S_i \cap [-m, m]) > 0$ . Choose  $n_0$  s.t.  $s - n_0\delta < -m < m < s + n_0\delta$ . Then  $P(s, \cdot)$  has positive density on  $[s - \delta, s + \delta]$  which implies  $P^{(n)}(s, \cdot)$  has positive density on  $[s - n\delta, s + n\delta] > 0$  for every  $n \geq n_0$  and in particular for  $n = n_0$  and  $n = n_0 + 1$ . But if  $b \geq 2$  it is not possible that  $b|n_0$  and  $b|n_0 + 1$  and so  $b = 1$ . ■

**Definition VI.14** A MC with state space  $\mathcal{S}$  has *stationary distribution*  $\pi$  if  $\Pi(C) = \int_{\mathcal{S}} P(s, C) \Pi(ds)$  for all  $C \in \mathcal{C}$ .

- as in the discrete case this implies that when  $\nu = \pi$  then  $X_n \sim \pi$  for all  $n$

- also if  $P(s, C)$  has density,  $f(x | s)$  then  $\pi$  is stationary when

$$\pi(x) = \int_{\mathcal{S}} f(x | s) \pi(s) ds$$

- if the Markov kernel  $P(s, \cdot)$  has density  $f(x | s)$  then it is *reversible wrt*  $\pi$  whenever  $f(x | s) \pi(s) = f(s | x) \pi(x)$  and then

$$\int_{\mathcal{S}} f(x | s) \pi(s) ds = \int_{\mathcal{S}} f(s | x) \pi(x) ds = \pi(x)$$

and so  $\pi$  is stationary



**Proposition VI.37** If a MC is  $\phi$ -irreducible, aperiodic and has stationary distribution  $\Pi$ , then

$$\lim_{n \rightarrow \infty} P(X_n \in C \mid X_0 = s) = \Pi(C)$$

for all  $s \in \mathcal{S}$  excepting possibly a set having  $\Pi$  measure 0.

Proof: fact

**Example VI.11** (*continued*)

- recall  $f(x \mid s)$  is the  $N(s/2, 1)$  density so  $f(s \mid x)$  is the  $N(x/2, 1)$  density
- so for reversibility we want

$$\begin{aligned} \exp(-(x - s/2)^2/2)\pi(s) &= \exp(-(x/2 - s)^2/2)\pi(x) \text{ or} \\ \frac{\pi(s)}{\pi(x)} &= \exp\left(-\frac{(x/2 - s)^2}{2} + \frac{(x - s/2)^2}{2}\right) \\ &= \exp\left(\frac{3}{8}(x^2 - s^2)\right) \end{aligned}$$

which is satisfied by  $\pi = N(0, 4/3)$

- also the chain satisfies Propositions VI.35 and VI.36 (normal densities are positive everywhere) so Prop. VI.37 applies

## Example VI.12 Markov Chain Monte Carlo

- suppose we want to sample from a distribution  $\pi$  and we can't find a good (e.g., rejection or inversion) algorithm to do this directly
- we try to create a MC, that can simulated from, s.t.  $\pi$  is its unique stationary distribution so that  $X_n$  is approximately distributed  $\pi$  for large  $n$
- let  $q(s, \cdot)$  be a proposal density for each  $s \in \mathcal{S}$  and suppose  $q$  is symmetric, i.e.,  $q(s, x) = q(x, s)$
- define  $P(s, \cdot)$  by

$$P(s, C \setminus \{s\}) = \int_{C \setminus \{s\}} q(s, x) \min \left( 1, \frac{\pi(x)}{\pi(s)} \right) dx$$

$$P(s, \{s\}) = 1 - \int_{\mathcal{S}} q(s, x) \min \left( 1, \frac{\pi(x)}{\pi(s)} \right) dx$$

and note

$$\int_{\mathcal{S}} q(s, x) \min \left( 1, \frac{\pi(x)}{\pi(s)} \right) dx \leq \int_{\mathcal{S}} q(s, x) dx = 1$$

so  $P(s, \cdot)$  is a probability measure

- the chain may stay at state  $s$  with possibly positive probability
- this is known as the *Metropolis algorithm*

1. given  $X_{n-1} = s$  generate  $Y_n \sim q(s, \cdot)$  stat. ind. of  $U_n \sim U(0, 1)$

2. put

$$X_n = \begin{cases} Y_n & \text{if } U_n \leq \pi(Y_n)/\pi(X_{n-1}) \\ X_{n-1} & \text{otherwise} \end{cases}$$

- to see that this works

$$\begin{aligned} & P(X_n \in C \setminus \{s\} \mid X_{n-1} = s) \\ = & P(Y_n \in C \setminus \{s\}, U_n \leq \pi(Y_n)/\pi(s) \mid X_{n-1} = s) \\ = & \int_{C \setminus \{s\}} \int_0^{\min(1, \pi(x)/\pi(s))} q(s, x) \, du \, dx \\ = & \int_{C \setminus \{s\}} q(s, x) \min(1, \pi(x)/\pi(s)) \, dx \end{aligned}$$

**Proposition VI.38** (*Continuous MCMC Convergence*) If  $\pi(s) > 0$  for every  $s \in \mathcal{S}$  and there is  $\delta > 0$  s.t.  $q(s, x) > 0$  whenever  $|x - s| < \delta$ , then  $X_n \xrightarrow{d} \pi$ .

Proof: We have, for  $s \neq x$

$$\begin{aligned} & \pi(s)q(s, x) \min(1, \pi(x)/\pi(s)) \\ = & \begin{cases} \pi(s)q(s, x) & \text{if } \pi(x)/\pi(s) > 1 \\ \pi(x)q(s, x) & \text{if } \pi(x)/\pi(s) \leq 1 \end{cases} \\ = & \pi(x)q(x, s) \min(1, \pi(s)/\pi(x)) \end{aligned}$$

and so the chain is reversible and  $\pi$  is a stationary distribution for the chain. Also,  $\pi(s)q(s, x) \min(1, \pi(x)/\pi(s)) > 0$  whenever  $|x - s| < \delta$  and so  $P(s, \cdot)$  has positive density when  $|x - s| < \delta$  and so the chain is  $\phi$ -irreducible ( $\phi = \text{volume}$ ) and aperiodic. The result then follows from Prop. VI.37. ■

**Exercise VI.18** Text 4.7.11

**Exercise VI.19** Text 4.8.3