Probability and Stochastic Processes II - Lecture 3b

Michael Evans University of Toronto https://utstat.utoronto.ca/mikevans/stac62/staC632024.html

2024

Michael Evans University of Toronto https://Probability and Stochastic Processes II - Lect

2024 1 / 16

III.2 Recurrence and Transience

- for MC $\{X_n : n \in \mathbb{N}_0\}$ with state space $S \subset \mathbb{Z}$ let $P_i = P(\cdot | X_0 = i)$ (so $p_{ii}^{(n)} = P_i(X_n = j)$) and let E_i denote expectation wrt P_i

Definition III.2.1 The (i, j)-th return probability for the MC $\{X_n : n \in \mathbb{N}_0\}$ is given by

$$f_{ij} = P_i(X_n = j \text{ for some } n) = \sum_{n=1}^{\infty} P_i(X_1 \neq j, \dots, X_{n-1} \neq j, X_n = j)$$

= the probability the chain visits state *i* after starting in state *i*

the probability the chain visits state j after starting in state i

- SO

 f_{ii} = the prob. the chain returns to state *i* after starting in state *i* $1 - f_{ii}$ = the prob. the chain never hits state *j* after starting in state *i* - note

$$P_i(X_n = j \text{ and } X_{n+m} = k \text{ for some } n \text{ and } m)$$

and by TTP
$$= \sum_{n=1}^{\infty} P(X_{n+m} = k \text{ for some } m \mid X_0 = i, X_1 \neq j, \dots, X_{n-1} \neq j, X_n = j) \times P_i(X_1 \neq j, \dots, X_{n-1} \neq j, X_n = j) \text{ and by Prop. III.3 and TH}$$

$$= \sum_{n=1}^{\infty} P(X_m = k \text{ for some } m \mid X_0 = j) P_i(X_1 \neq j, \dots, X_{n-1} \neq j, X_n = j)$$

$$= f_{jk} \sum_{n=1}^{\infty} P_i(X_1 \neq j, \dots, X_{n-1} \neq j, X_n = j) = f_{ij}f_{jk}$$

3

イロト イヨト イヨト イヨト

- let $N_i = \#\{n : X_n = i\} \in \mathbb{N}_0 \cup \{\infty\}$ and N_i is the number of visits the chain makes to state i

- therefore the conditional distribution of N_j given $X_0 = i$ is given by (using argument on previous slide)

$$P_{i}(N_{j} = 0) = 1 - f_{ij}$$

$$P_{i}(N_{j} \ge 1) = f_{ij}, P_{i}(N_{j} \ge 2) = f_{ij}f_{jj}, \dots, P_{i}(N_{j} \ge k) = f_{ij}f_{jj}^{k-1}$$

$$P_{i}(N_{j} = k) = P_{i}(N_{j} \ge k) - P_{i}(N_{j} \ge k+1) = f_{ij}f_{jj}^{k-1}(1 - f_{jj})$$

$$P_{i}(N_{i} = k) = f_{ii}^{k}(1 - f_{ii}) \ (= 0 \text{ when } f_{ii} = 1)$$

$$P_{i}(N_{i} = \infty) = 1 - \sum_{k=0}^{\infty} P_{i}(N_{i} = k) = \begin{cases} 1 & f_{ii} = 1 \\ 0 & f_{ii} < 1 \end{cases}$$

Definition III.2.1 A state *i* of a MC is *recurrent* if $f_{ii} = 1$ and is *transient* if $f_{ii} < 1$.

Proposition III.6 (Recurrent state theorem) For a MC

(i) state *i* is recurrent iff $P_i(N_i = \infty) = 1$ iff $E_i(N_i) = \sum_{n=1}^{\infty} p_{ii}^{(n)} = \infty$, (ii) state *i* is transient iff $P_i(N_i = \infty) = 0$ iff $E_i(N_i) = \sum_{n=1}^{\infty} p_{ii}^{(n)} < \infty$.

Proof: The first parts of (i) and (ii) follow from the conditional distribution of N_i given $X_0 = i$ just worked out. Now note that $N_i = \sum_{n=1}^{\infty} I_{\{i\}}(X_n)$ and so

$$E_{i}(N_{i}) = E_{i}\left(\sum_{n=1}^{\infty} I_{\{i\}}(X_{n})\right) \stackrel{MCT}{=} \sum_{n=1}^{\infty} E_{i}(I_{\{i\}}(X_{n})) = \sum_{n=1}^{\infty} p_{ii}^{(n)}$$

and when *i* is recurrent $P_i(N_i = \infty) = 1$ so $E_i(N_i) = \infty$. Now when $f_{ii} < 1$, then

$$E_{i}(N_{i}) = \sum_{k=1}^{\infty} k f_{ii}^{k} (1 - f_{ii}) = (f_{ii} - f_{ii}^{2}) + 2(f_{ii}^{2} - f_{ii}^{3}) + \dots$$
$$= \sum_{k=1}^{\infty} f_{ii}^{k} = \left(\frac{1}{1 - f_{ii}} - 1\right) = \frac{f_{ii}}{1 - f_{ii}} < \infty.$$
So when $\sum_{n=1}^{\infty} p_{ii}^{(n)} = \infty$ then state *i* is recurrent and is otherwise

Michael Evans University of Toronto https://Probability and Stochastic Processes II - Lect

Side note - an alternative expression for the expectation of a r. v.

Lemma If
$$X \ge 0$$
 then $E(X) = \int_0^\infty (1 - F_X(z)) dz$ and when $P(X \in \mathbb{N}_0) = 1$ this equals $E(X) = \sum_{k=1}^\infty P(X \ge k)$.

Proof: We have

$$E(X) = \int_0^\infty x P_X(dx) = \int_0^\infty \int_0^x dz \, P_X(dx)$$

=
$$\int_0^\infty \int_0^\infty I_{(0,x)}(z) dz \, P_X(dx) \stackrel{Fubini}{=} \int_0^\infty \int_0^\infty I_{(0,x)}(z) P_X(dx) \, dz$$

=
$$\int_0^\infty P(X > z) \, dz = \int_0^\infty (1 - F_X(z)) \, dz$$

=
$$\sum_{k=1}^\infty P(X \ge k) \text{ when } P(X \in \mathbb{N}_0) = 1. \blacksquare$$

Exercise III.2.1 If E(X) exists, then prove

$$E(X) = \int_0^\infty (1 - F_X(z)) \, dz - \int_{-\infty}^0 F_X(z) \, dz$$

= $\sum_{k=1}^\infty P(X \ge k) - \sum_{k=-1}^\infty P(X \le k)$ when $P(X \in \mathbb{N}_0) = 1$.

Michael Evans University of Toronto https://Probability and Stochastic Processes II - Lect

Example III.2.1 Simple random walk

- we showed (assuming $X_0 = 0$)

$$p_{ij}^{(n)} = \begin{cases} 0 & \text{if } n+j-i \text{ not even} \\ \left(\frac{n}{n+j-i}\right)p^{\frac{n+j-i}{2}}(1-p)^{n-\frac{n+j-i}{2}} & \text{if } n+j-i \in \{0, 2, \dots, 2n\} \end{cases}$$

- therefore

$$\sum_{n=1}^{\infty} p_{ii}^{(n)} = \sum_{n=1}^{\infty} {\binom{2n}{n}} p^n \left(1-p\right)^n = \sum_{n=1}^{\infty} \frac{(2n)!}{n!n!} p^n \left(1-p\right)^n$$

- Stirling's approximation to *n*! says, putting $s_n = (n/e)^n \sqrt{2\pi n}$, that $\lim_{n\to\infty} n!/s_n = 1$ and so

$$\lim_{n \to \infty} \frac{(2n)!}{n!n!} \frac{s_n^2}{s_{2n}} = 1 \text{ and}$$
$$t_n = \frac{s_{2n}}{s_n^2} = \frac{2^{2n} (n/e)^{2n} \sqrt{4\pi n}}{(n/e)^{2n} (2\pi n)} = \frac{2^{2n}}{\sqrt{\pi n}}$$

- now

$$\sum_{n=1}^{\infty} \frac{(2n)!}{n!n!} p^n (1-p)^n = \sum_{n=1}^{\infty} \left(\frac{(2n)!}{n!n!t_n}\right) t_n p^n (1-p)^n$$
$$= \sum_{n=1}^{\infty} \left(\frac{(2n)!}{n!n!t_n}\right) \frac{(4p(1-p))^n}{\sqrt{\pi n}}$$

and note that $0 \leq 4p(1-p) \leq 1$ is maximized at p=1/2 and equals 1 there

- for large *n* the *n*-th term in this series is like $\frac{(4p(1-p))^n}{\sqrt{\pi n}}$ and

$$\sum_{n=1}^{\infty} \frac{(4p(1-p))^n}{\sqrt{\pi n}} \begin{cases} < \infty & p \neq 1/2 \\ = \infty & p = 1/2 \end{cases}$$

and

$$0 < b \le \left(\frac{(2n)!}{n!n!t_n}\right) \le B \text{ for all } n \text{ for some } (b, B) \text{ which implies}$$
$$\sum_{n=1}^{\infty} p_{ii}^{(n)} = \begin{cases} < \infty & p \ne 1/2 \\ = \infty & p = 1/2 \end{cases}$$

and so any i is recurrent for a symmetric srw and in the nonsymmetric case all states i are transient \blacksquare

Exercises

III.2.2 Text 1.5.12

III.2.3 Text 1.5.15

III.3 Communicating States and Irreducibility

Definition III.3.1 State *i* communicates with state *j* if $f_{ij} > 0$ denoted by $i \rightarrow j$ (equivalently there exist *n* s.t. $p_{ij}^{(n)} > 0$). A MC is *irreducible* if $i \rightarrow j$ for all $i, j \in S$.

Lemma III.7 (Sum Lemma) If $i \to k$ and $l \to j$ and $\sum_{n=1}^{\infty} p_{kl}^{(n)} = \infty$, then $\sum_{n=1}^{\infty} p_{ij}^{(n)} = \infty$. In particular, if $i \leftrightarrow k$ and $\sum_{n=1}^{\infty} p_{kk}^{(n)} = \infty$, then $\sum_{n=1}^{\infty} p_{ii}^{(n)} = \infty$ so if i and k mutually communicate and one is recurrent then both are recurrent.

Proof: There exist $m, r \ge 1$ s.t. $p_{ik}^{(m)} > 0, p_{lj}^{(r)} > 0$. The Chapman-Kolmogorov inequality implies $p_{ij}^{(m+s+r)} \ge p_{ik}^{(m)} p_{kl}^{(s)} p_{lj}^{(r)}$ for all s. Therefore,

$$\sum_{n=1}^{\infty} p_{ij}^{(n)} \ge \sum_{s=1}^{\infty} p_{ij}^{(m+s+r)} \ge \sum_{s=1}^{\infty} p_{ik}^{(m)} p_{kl}^{(s)} p_{lj}^{(r)} = p_{ik}^{(m)} p_{lj}^{(r)} \sum_{s=1}^{\infty} p_{kl}^{(s)} = \infty$$

as required. Now suppose $i \leftrightarrow k$ and put j = i, l = k which implies $\sum_{n=1}^{\infty} p_{ii}^{(n)} = \infty$ whenever $\sum_{n=1}^{\infty} p_{kk}^{(n)} = \infty$ and conversely.

Proposition III.8 (*Cases theorem*) For an irreducible MC either (i) $\sum_{n=1}^{\infty} p_{ij}^{(n)} = \infty$ for all *i*, *j* and so all states are recurrent or (ii) $\sum_{n=1}^{\infty} p_{ij}^{(n)} < \infty$ for all *i*, *j* and all states are transient.

Proof: (i) If $\sum_{n=1}^{\infty} p_{kl}^{(n)} = \infty$ for some k, l, then by Lemma III.7 $\sum_{n=1}^{\infty} p_{ij}^{(n)} = \infty$ for all i, j which implies $\sum_{n=1}^{\infty} p_{ii}^{(n)} = \infty$ for all i. (ii) If $\sum_{n=1}^{\infty} p_{kl}^{(n)} < \infty$ for all k, l then $\sum_{n=1}^{\infty} p_{ii}^{(n)} < \infty$ for all i.

Example III.3.1 Simple random walk

- we have

$$p_{ij}^{(n)} = \begin{cases} 0 & \text{if } n+j-i \text{ not even} \\ \left(\frac{n}{n+j-i}\right)p^{\frac{n+j-i}{2}}(1-p)^{n-\frac{n+j-i}{2}} & \text{if } n+j-i \in \{0, 2, \dots, 2n\} \end{cases}$$

so choosing *n* s.t. n + j - i is even we have that $p_{ij}^{(n)} > 0$ so $i \to j$ and similarly $j \to i$ and a srw is irreducible \blacksquare

Proposition III.9 (*Finite State Space Theorem*) If the MC is irreducible and *S* is finite then the chain is recurrent.

Proof: We have

$$\sum_{j\in\mathcal{S}}\sum_{n=1}^{\infty} p_{ij}^{(n)} \stackrel{ extsf{Fubini}}{=} \sum_{n=1}^{\infty}\sum_{j\in\mathcal{S}} p_{ij}^{(n)} = \sum_{n=1}^{\infty} 1 = \infty$$

and since $\sum_{j \in S}$ is a finite sum this implies $\sum_{n=1}^{\infty} p_{ij}^{(n)} = \infty$ for some j and result follows.

2024 12 / 16

- put

Lemma III.10 (*Hit Lemma*) If $j \rightarrow i$ with $j \neq i$ then $P_j(H_{ij}) > 0$.

Proof: Since $j \to i$ there exists a sequence of states i_0, i_1, \ldots, i_m with $i_0 = j, i_m = i$ and such that $p_{i_0i_1}p_{i_1i_2} \ldots p_{i_{m-1}i_m} > 0$. Then there is a shortest such subsequence within this sequence which starts with j, and since this chain is in H_{ji} this proves the result.

Lemma III.11 (*f-Lemma*) If $j \rightarrow i$ and $f_{jj} = 1$, then $f_{ij} = 1$.

Proof: This is clear when i = j so assume $i \neq j$. Then $P_j(H_{ij}) > 0$. Now

$$0 = 1 - f_{jj} = P_j (\text{never return to } j)$$

$$\geq P_j (\text{for some } n, X_1 \neq j, \dots X_{n-1} \neq j, X_n = i \text{ and } X_{n+m} \neq j \forall m \ge 1)$$

$$= \sum_{n=1}^{\infty} P_j (X_1 \neq j, \dots X_{n-1} \neq j, X_n = i \text{ and } X_{n+m} \neq j \forall m \ge 1)$$

$$= \sum_{n=1}^{\infty} P_j (X_1 \neq j, \dots X_{n-1} \neq j, X_n = i) P(X_{n+m} \neq j \forall m \ge 1 | X_n = i)$$

$$= \sum_{n=1}^{\infty} P_j (X_1 \neq j, \dots X_{n-1} \neq j, X_n = i) P_i (X_m \neq j \forall m \ge 1) (\text{TH})$$

$$= (1 - f_{ij}) \sum_{n=1}^{\infty} P_j (X_1 \neq j, \dots X_{n-1} \neq j, X_n = i) = (1 - f_{ij}) P_j (H_{ij}) \ge 0$$

and this implies $f_{ij} = 1$.

024 14 / 16

Lemma III.12 (Infinite Returns Lemma) For an irreducible MC, $P_i(N_j = \infty) = \begin{cases} 1 & \text{for every } i, j \text{ if chain is recurrent} \\ 0 & \text{for every } i, j \text{ if chain is transient.} \end{cases}$ Proof: This follows from the earlier derivation

$$P_i(N_j = k) = f_{ij}f_{jj}^{k-1}(1 - f_{jj})$$

since this implies $P_i(N_j = k) = 0$ for every $k \in \mathbb{N}$ when the chain is recurrent since $f_{jj} = 1$. When $f_{jj} < 1$, so the chain is transient, then

$$P_i(N_j = \infty) = \lim_{k \to \infty} P_i(N_j \ge k) = \lim_{k \to \infty} f_{ij} f_{jj}^{k-1} = 0.$$

- so there are a number of ways to characterize the recurrence or transience of an irreducible MC and the book presents a nice summary (pages 22-23.)

- a subset $C \subset S$ is called *closed* if for every $i \in C, j \notin C$ then $f_{ij} = 0$ so once we get into C we can never leave

- so we can consider a new chain with state space C once we enter C which may be recurrent or transient.

Exercises

- III.3.1 Text 1.6.4
- III.3.2 Text 1.6.20
- III.3.2 Text 1.6.24

э

(日) (同) (三) (三)